ARTICLE IN PRESS

RESUSCITATION XXX (2020) XXX -XXX



Available online at www.sciencedirect.com

Resuscitation



journal homepage: www.elsevier.com/locate/resuscitation

¹ Clinical paper

² Randomized trial of the i-gel supraglottic airway

- device versus tracheal intubation during out of
- 4 hospital cardiac arrest (AIRWAYS-2): Patient
- outcomes at three and six months
- ⁶Q1 Jonathan R. Benger^{a, *}, Michelle J. Lazaroo^b, Madeleine Clout^b, Sarah Voss^a,
- ⁷ Sarah Black^c, Stephen J. Brett^d, Kim Kirby^{a,c}, Jerry P. Nolan^{e, f}, Barnaby C. Reeves^b,
- ⁸ Maria Robinson^c, Lauren J. Scott^{b,g}, Helena Smartt^b, Adrian South^c, Jodi Taylor^{b,e},
- ⁹ Matthew Thomas^{*i*}, Sarah Wordsworth^{*h*}, Chris A. Rogers^{*b*}
- ¹⁰ ^a University of the West of England, Glenside Campus, Bristol, UK
- ¹¹ ^b Clinical Trials and Evaluation Unit (CTEU), Bristol Trials Centre, Bristol Medical School, University of Bristol, Bristol, UK
- ¹² ^c South Western Ambulance Service NHS Foundation Trust, Exeter, UK
- ¹³ ^d Department of Surgery and Cancer, Imperial College, London, UK
- ¹⁴ ^e Bristol Medical School, University of Bristol, Bristol, UK
- ¹⁵ ^f Department of Anaesthesia, Royal United Hospital, Bath, UK
- ¹⁶ ⁹ National Institute for Health Research Applied Research Collaboration West (NIHR ARC West), University Hospitals Bristol and Weston NHS
- ¹⁷ Foundation Trust, Bristol, UK
- ¹⁸ ^h Health Economics Research Centre, Nuffield Department of Population Health, University of Oxford, Oxford, UK
- ¹⁹ ⁱ Intensive Care Unit, University Hospitals Bristol NHS Foundation Trust, Bristol, UK

Abstract

20

Aim: The AIRWAYS-2 cluster randomised controlled trial compared the i-gel supraglottic airway device (SGA) with tracheal intubation (TI) as the first advanced airway management (AAM) strategy used by Emergency Medical Service clinicians (paramedics) treating adult patients with non-traumatic out-of-hospital cardiac arrest (OHCA). It showed no difference between the two groups in the primary outcome of modified Rankin Scale (mRS) score at 30 days/hospital discharge. This paper reports outcomes to 6 months.

Methods: Paramedics from four ambulance services in England were randomised 1:1 to use an i-gel SGA (759 paramedics) or TI (764 paramedics) as their initial approach to AAM. Adults who had a non-traumatic OHCA and were attended by a participating paramedic were enrolled automatically under a waiver of consent. Survivors were invited to complete questionnaires at three and six months after OHCA. Outcomes were analysed using regression methods. **Results:** 767/9296 (8.3%) enrolled patients survived to 30 days/hospital discharge and 317/767 survivors (41.3%) consented and were followed-up to six months. No significant differences were found between the two treatment groups in the primary outcome measure (mRS score: 3 months: odds ratio (OR) for good recovery (i-gel/TI, OR) 0.89, 95% CI 0.69–1.14; 6 months OR 0.91, 95% CI 0.71–1.16). EQ-5D-5L scores were also similar between groups and sensitivity analyses did not alter the findings.

Conclusion: There were no statistically significant differences between the TI and i-gel groups at three and six months. We therefore conclude that the initially reported finding of no significant difference between groups at 30 days/hospital discharge was sustained when the period of follow-up was extended to six months.

Keywords: Heart arrest, Airway management, Tracheal intubation, Laryngeal mask, Survival rate, Health-related quality of life

Received 15 August 2020; Accepted 21 September 2020 Available online xxx

0300-9572/© 2020 The Author(s). Published by Elsevier B.V. This is an open access article under the CC BY-NC-ND license (http://creativecommons.org/

^{*} Corresponding author at: Faculty of Health and Life Sciences, University of the West of England, Glenside Campus, Bristol BS16 1DD, UK. E-mail address: Jonathan.Benger@uwe.ac.uk (J.R. Benger).

https://doi.org/10.1016/j.resuscitation.2020.09.026

RESUS 8713 1-8

2

RESUSCITATION XXX (2020) XXX -XXX

21 Introduction

2203 Survival rates following out-of-hospital cardiac arrest (OHCA) remain 23 disappointingly low. Of the nearly 30,000 people who receive 24 resuscitation for OHCA in England annually, only 25% achieve a 25 return of spontaneous circulation (ROSC), and 8% are discharged 26 from hospital alive.¹

27 The earlier an intervention is provided in OHCA the greater its 28 potential to increase survival.² If basic life support and initial 29 defibrillation of a shockable rhythm does not result in ROSC, the 30 attention of emergency medical services (EMS) clinicians (para-31 medics) turns to airway management and drug delivery.³ 32 However, optimal airway management during OHCA has been 33 an enduring area of uncertainty, with very little high-quality 34 research on which to base treatment recommendations.⁴ Options 35 range from basic airway intervention to advanced procedures such 36 as the insertion of a supraglottic airway (SGA) or tracheal 37 intubation (TI), which is considered the "gold standard" of 38 advanced airway management.⁵

39 Large observational studies have shown an association between 40 survival following OHCA and the use of basic airway management 41 techniques, when compared with either SGA or TI.⁶ However these 42 studies are prone to residual confounding and resuscitation time 43 bias.^{7,8} As a result, methodologies to complete high-quality random-44 ised controlled trials (RCTs) of alternative advanced airway 45 management (AAM) strategies in the early stages of cardiac arrest 46 have been developed.9

47 During 2018, two RCTs of AAM during OHCA were published. 48 Both compared an SGA with TI as the initial AAM strategy adopted 49 by paramedics treating non-traumatic OHCA in adults. The 50 Pragmatic Airway Resuscitation Trial (PART) compared the 51 laryngeal tube SGA with TI in 3004 patients and found a statistically 52 significant benefit in survival to 72h and hospital discharge, and a 53 favourable neurological status at hospital discharge, for those 54 patients randomised to the laryngeal tube.¹⁰ At the same time, we 55 published the AIRWAYS-2 trial which showed no difference in good 56 functional outcome (modified Rankin Scale (mRS) score of 0-3) at 57 hospital discharge or 30 days between 9296 patients randomised to 58 either the i-gel SGA or TI.11

59 Whilst early outcomes are valuable measures in studies involving 60 OHCA patients, there is an increasing recognition of the importance of 61 longer-term outcomes and functional recovery following OHCA, 62 including quality of life in survivors.¹² The aim of this paper was 63 therefore to compare the secondary outcomes of mRS score and 64 EuroQol-5D (EQ-5D) at three and six months after OHCA between 65 groups of patients in the AIRWAYS-2 trial managed by paramedics 66 randomised to use either the i-gel or TI as their initial AAM strategy 67 when treating adult patients following OHCA.

68 **Methods**

69 The AIRWAYS-2 trial methodology has been reported previously.^{11,13} 70 In summary, we completed a cluster RCT of paramedics from four 71 large EMS provider organisations (ambulance services) in England 72 covering approximately 21 million people. 1523 paramedics volun-73 teered to participate and were randomised 1:1 to use an i-gel SGA 74 (759 paramedics) or TI (764 paramedics) as their initial AAM strategy 75 when attending adult patients with non-traumatic OHCA.

Randomisation and case ascertainment

77 Individual patient randomisation was considered impractical due to the risk that research procedures would delay life-saving treatment. 79 We therefore chose to designate paramedics as the unit of random-80 isation, thereby creating a relatively large number of clusters 81 containing a relatively small number of patients on average. This 82 had the benefit of minimising the intra-cluster correlation and more closely approximating individual patient randomisation than would be the case if larger clusters were used. It was not possible to blind paramedics to the treatment allocation. Therefore, it was necessary to enrol all eligible patients into the trial to avoid the risk of selection bias, 87 (e.g. to avoid paramedics selectively enrolling patients on the basis of 88 their predicted outcome).¹⁴ This complete case ascertainment was achieved by supplementing routine case reporting by participating 90 paramedics with daily review of all cardiac arrests occurring in the four 91 participating EMS provider organisations (ambulance services) and 92 cross-referencing with routinely collected audit data that are submitted 93 to a national OHCA registry.¹

76

78

83

84

85

86

89

94

96

100

101

102

103

104

105

106

107

108

109

122

Patient enrolment

95 Automatic patient enrolment proceeded under a waiver of consent provided by the Confidentiality Advisory Group (CAG: reference 14/ 97 CAG/1030). Ethics review and approval was provided by South 98 Central - Oxford C Research Ethics Committee (REC: reference 14/ 99 SC/1219). This included a process of written informed consent for participating paramedics and for surviving patients (or a personal consultee for surviving patients without mental capacity). The main disadvantage of automatic enrolment was that many enrolled patients did not receive any AAM. There was also an increased risk that eligible patients might not be recognised as such by the participating paramedic, leading to protocol deviations. Paramedics were given the clinical freedom to deviate from the trial protocol if they felt that a particular approach to airway management was in the patient's best interests.

Patient inclusion criteria

110 Patient inclusion criteria were: known or believed to be 18 years of age 111 or older; non-traumatic OHCA; attended by a paramedic participating 112 in the trial who was either the first or second paramedic to arrive at the 113 patient's side; resuscitation commenced or continued by paramedics 114 or EMS personnel. Patient exclusion criteria were: detained in the 115 Prison Service; previously recruited to the trial (determined retro-116 spectively); resuscitation deemed inappropriate (using guidelines 117 based on those of the Joint Royal Colleges Ambulance Liaison 118 Committee)¹⁵; advanced airway already in place (inserted by another 119 paramedic, doctor or nurse) when a paramedic participating in the trial 120 arrived at the patient's side; known to be already enrolled in another 121 pre-hospital RCT; patient mouth opening <2cm.

Intervention

123 The intervention was the insertion of a second generation SGA (i-gel: 124 Intersurgical, Wokingham, UK), which is the SGA most commonly 125 used by paramedics in England.¹⁶ This was compared with TI using 126 direct larvngoscopy and an intubating bougie. A standard approach to 127 airway management in the trial, from basic to advanced techniques, 128 was agreed by participating ambulance services. This included the

ARTICLE IN PRESS

RESUSCITATION XXX (2020) XXX -XXX

154

use of bag-mask ventilation and simple airway adjuncts before AAM.
 In all other respects care proceeded as usual, with resuscitation
 following standard international guidelines.¹⁷

¹³² Follow-up procedures

133 All enrolled patients who survived to hospital discharge were 134 followed-up by a member of the local research team who 135 consulted with the clinical staff caring for the patient to determine 136 the optimal time to approach the patient and/or their family to seek 137 consent. Consent was sought from the patient or from a personal 138 consultee if the patient was judged to lack capacity. Each patient 139 or consultee was able to choose either active follow-up (collection 140 of routinely available data combined with telephone and/or postal 141 contact at 30 days/hospital discharge, and three and six months 142 after cardiac arrest), passive follow-up (collection of routinely 143 available data only, with no further patient contact) or no further 144 data collection. In cases where a consultee opinion was obtained, 145 and active follow-up chosen, the patient's capacity was re-146 assessed at the three- and six-month follow-up. If an initially 147 incapacitated patient regained capacity, consent to continue their 148 involvement in the trial was sought from the patient. The mortality 149 status of patients who consented to follow-up was ascertained 150 from national record systems. The mortality status of all other 151 survivors was obtained from Hospital Episode Statistics data 152 provided by NHS Digital (under HRA CAG approval) where 153 linkage was possible.

Outcomes

155 For patients who provided active consent, the mRS score was 156 measured at three and six months after the index OHCA. The mRS 157 score, which incorporates both functional outcome and survival, is 158 widely used in OHCA research and comprises a seven-point scale (0 159 -6) with lower scores representing better recovery.¹⁸ Patients who die 160 are given a score of six. The mRS scores were dichotomised into good 161 recovery (score 0-3) and poor recovery (score 4-6). The EuroQol 162 (EQ-5D-5L) is a validated measure of health-related guality of life and has been used widely in OHCA survivors.¹⁹ The EQ-5D-5L descriptive 163 164 system comprises 5 domains: mobility, self-care, usual activities, pain/ 165 discomfort, and anxiety/depression. The EQ-5D-5L visual analogue 166 scale (VAS) records a person's self-rated health with a range of 0 167 -100. The EQ-5D-5L descriptive system and VAS were measured at 168 30 days/hospital discharge (whichever comes first), three and six 169 months after the index OHCA for patients who consented to active 170 follow-up and had survived to these timepoints. The EQ-5D-5L index 171 scores (index) were calculated from the descriptive system responses 172 by mapping onto the EQ5D-3L value set.²⁰ Patients who had died 173 were given a value of 0 for both the index and VAS scores.

Statistical analysis

174

The primary analysis included participants with outcome data (complete-case analysis). The effect of missing data was examined with two sensitivity analyses. The first ('worst-case' scenario) 177



Fig. 1 - Flow of participants and data.

^[1] 10 patients (5 TI, 5 i-gel) who did not consent to follow-up have unknown survival status.

^[2] 1 patient (1 TI, 0 i-gel) who consented to active follow-up is known to have survived to 3 months but has unknown survival status between 3 months and 6 months follow-up.

4

ARTICLE IN PRESS

RESUSCITATION XXX (2020) XXX -XXX

178 assigned the worst possible score to known survivors with missing 179 data, whilst patients for whom the survival status was unknown were 180 assumed to have died. The second sensitivity analysis ('imputed case' 181 scenario), used multiple imputation (60 imputations). Imputations was 182 performed using the ICE command in Stata v15.1 (StataCorp) and 183 estimates were combined using Rubin's rules. The imputation model 184 included the following variables: age, sex, length of intensive care unit 185 (ICU) stay, treatment group, ambulance provider organisation, 186 paramedic experience, distance from base ambulance station, and 187 index, VAS and mRS scores at 30 days/hospital discharge, three 188 months and six months timepoints (see supplement for further 189 details).

Logistic regression was used to analyse the dichotomised mRS
 scores with paramedic fitted as a random effect. A two-part binomial beta model was used to analyse the EQ-5D-5L index and VAS scores.

The scores of survivors were transformed for the purposes of 193 modelling using the following transformation: 194

$$y'=rac{y-a}{b-a}$$

$$y^n = \frac{[y'(N-1)+1/2]}{N}$$

195

196

197

198

199

200

where *y* is the outcome (index or VAS score), *b* is the highest possible score (index: 1, VAS: 100), *a* is the smallest possible score (index: -0.59, VAS: 0), *N* is the total number of survivors with data and y^n is the transformed score. This transformation ensured that the transformed scores were between 0 and 1 (excluding 0 and 1) which is

Table 1 - Complete case modified Rankin Scale analyses results.

Complete case modified Rankin Scale (0-3; good recovery)	Randomised to Tracheal Intubation (<i>n</i> =4410)	Randomised to i- gel (n=4886)	Odds Ratio estimate (95% CI)	<i>p</i> - Value	ICC	Risk difference estimate (95% CI)	<i>p</i> - Value
	n (%)	n (%)					
Hospital discharge/30 days (mRS 0-3; good recovery) ^a	300/4407 (6.8%)	311/4882 (6.4%)	OR=0.92 (0.77, 1.09)	0.33	0.05	RD=-0.62% (-1.65%, +0.41%)	0.24
0 (no symptoms)	124/4407 (2.8%)	117/4882 (2.4%)					
1	48/4407 (1.1%)	41/4882 (0.8%)					
2	50/4407 (1.1%)	58/4882 (1.2%)					
3	78/4407 (1.8%)	95/4882 (1.9%)					
4	46/4407 (1.0%)	45/4882 (0.9%)					
5	27/4407 (0.6%)	39/4882 (0.8%)					
6 (deceased)	4034/4407 (91.5%)	4487/4882 (91.9%)					
Three months follow-up (mRS 0–3; good recovery) ^{a,b}	123/4199 (2.9%)	121/4636 (2.6%)	OR=0.89 (0.69, 1.14)	0.35	<0.001	RD=-0.51% (-1.18%, +0.16%)	0.14
0 (no symptoms)	52/4199 (1.2%)	55/4636 (1.2%)				+0.10 %)	
1	6/4199 (0.1%)	4/4636 (0.1%)					
2	30/4199 (0.7%)	35/4636 (0.8%)					
3	35/4199 (0.8%)	27/4636 (0.6%)					
4	22/4199 (0.5%)	17/4636 (0.4%)					
5	5/4199 (0.1%)	4/4636 (0.1%)					
6 (deceased)	4049/4199 (96.4%)	4494/4636 (96.9%)					
Non-active consent patients who were not	164/4407 (4.7%)	186/4882 (5.0%)					
known to have died at three months		(010,00)					
Six months follow-up (mRS 0–3; good recovery) ^{a,c}	134/4212 (3.2%)	136/4661 (2.9%)	0.91 (0.71, 1.16)	0.43	<0.001	RD=-0.39% (-1.08%, +0.30%)	0.27
0 (no symptoms)	59/4212 (1.4%)	66/4661 (1.4%)					
1	4/4212 (0.1%)	5/4661 (0.1%)					
2	42/4212 (1.0%)	41/4661 (0.9%)					
3	29/4212 (0.7%)	24/4661 (0.5%)					
4	18/4212 (0.4%)	18/4661 (0.4%)					
5	2/4212 (0.1%)	3/4661 (0.1%)					
6 (deceased)	4058/4212 (96.3%)	4504/4661 (96.6%)					
Non-active consent patients who were not	158/4407 (4.4%)	180/4882 (4.5%)					
known to have died at six months							

^a 7 patients (3 Tracheal Intubation, 4 i-gel) were unable to be identified at 30 days/hospital discharge and were excluded from this analysis.

^b 104 patients (44 Tracheal Intubation, 60 i-gel) were missing mRS at 3 months follow-up and were excluded from the complete-case analysis at this timepoint. ^c 78 patients (37 Tracheal Intubation, 41 i-gel) were missing mRS at 6 months follow-up and were excluded from the complete-case analysis at this timepoint.

RESUSCITATION XXX (2020) XXX

240

241

242

243

244

245

246

247

248

249

250

251

252

253

254

255

256

257

258

259

201 required for beta regression. The two-part binomial-beta model 202 produces two treatment estimates.²¹ The first (binomial part) is the 203 odds ratio for survival ('alive vs dead') with an estimate greater than 204 1 favouring i-gel over TI. The second estimate (beta part) relates to the 205 quality of life of survivors ('given patient survived'). Again, an estimate 206 greater than 1 indicates a better guality of life in the i-gel group over the 207 TI aroup.

208 All models were fitted to each timepoint separately as convergence 209 issues prevented the fitting of longitudinal models. To allow for 210 clustering of paramedics in the two-part binomial-beta models, 211 confidence intervals were estimated using clustered bootstrapping 212 (see supplement for further details). The clustered bootstrap and two-213 part binomial-beta model was performed in SAS v9.4. All other 214 analyses were performed in Stata v15.1 (StataCorp).

215 **Results**

216 In total, 9296 patients were enrolled in the AIRWAYS-2 trial (4410 TI, 217 4886 i-gel). 767/9296 (8.3%) of patients survived to 30 days/hospital 218 discharge and 402/767 (52.5%) consented to active follow-up. Of the 219 402 patients who consented to active follow-up, 388 (96.5%) were 220 known to have survived to six months post-OHCA, 13 had died and the 221 survival status at six months was unknown for 1 patient. All 222 402 patients who consented to active follow-up completed ques-223 tionnaires at 30 days/hospital discharge. Completion rates at three 224 months and six months were 300/396 (153/194 TI, 147/202 i-gel), and 225 317/388 (159/190 TI, 158/198 i-gel) respectively (Fig. 1).

226 In the period between 30 days/hospital discharge and six months 227 post-OHCA, the proportion of patients with a mRS score of 0 (no 228 symptoms) increased whilst the proportion of patients with a mRS 229 score of 5 (severe disability requiring constant nursing care) 230 decreased (Table 1, Supplement Fig. 1). All patients with a score 231 of 5 at six months also had a score of 5 at 30 days/hospital discharge 232 (Supplement Fig. 2). Most patients with a mRS score of 0 at three and 233 six months had improved since 30 days/hospital discharge

234 (Supplement Figs. 2 and 3). Of patients who had a mRS score of 235 0 at six months, the majority also had a score of 0 at three months 236 (Supplement Fig. 3). Of the 66 patients with a mRS score of 5 at 237 30 days/hospital discharge, 12/66 (18.2%) died before 3 months, 11/ 238 66 (16.7%) improved and 5/66 (7.6%) stayed the same; data were 239 missing for the remaining 38 (57.6%) (Supplement Fig. 4).

The mRS scores at all three timepoints showed higher proportions of patients with a good recovery in the TI group compared to the i-gel group but these differences were not statistically significant [complete case analysis: 30 days/hospital discharge OR=0.92 (95% CI 0.77-1.09); three months OR=0.89 (95% CI 0.69-1.14); six months OR=0.91 (95% CI 0.71-1.16)] (Table 1, Fig. 2)]. The 'worst-case'' and 'imputed case' sensitivity analyses provided consistent results (Supplement Tables 1 and 2, Fig. 2).

The EQ-5D domain scores are shown in Supplement Table 3. The data indicate higher median index and VAS scores at 30 days/hospital discharge in the TI group and similar median scores at the later timepoints (Table 2). The survival component of the two-part model showed no statistically significant difference in the odds of survival in the TI group compared to the i-gel group at all three timepoints (Table 2, Figs. 3 and 4). For the quality of life component in survivors, the outcomes were similar in the two groups at all timepoints for both index and VAS scores (Table 2, Figs. 3 and 4). The sensitivity analyses showed consistent findings with the complete case analyses (Supplement Tables 4 and 5, Supplement Figs. 5-8).

Discussion

260 The functional outcomes (mRS scores) at 3 and 6 months for patients 261 recruited to the AIRWAYS-2 trial were consistent with the primary 262 outcome of mRS score measured at 30 days/hospital discharge.¹¹ 263 The proportions of patients achieving a good recovery were not 264 statistically different between the two treatment groups at all three 265 timepoints. Quality of life measured using the EQ-5D-5L also revealed 266 no significant differences between the two treatment groups across



Modified Rankin Scale

Fig. 2 - Main analyses of modified Rankin Scale scores.

6

ARTICLE IN PRESS

R E S U S C I T A T I O N X X X (2020) X X X - X X X

Table 2 - Complete case EQ-5D-5L index and visual analogue scale analyses results.

	Randomised to Tracheal Intubation (<i>N</i> =4410)		Randomised to i-gel (<i>N</i> =4886)		'Alive vs dead' model		'Given patient survived' model	
	Survived (<i>n</i> (%))	Median (IQR)	Survived (<i>n</i> (%))	Median (IQR)	OR ^a (95% CI)	<i>p</i> - Value	OR ^b (95% CI)	<i>p</i> - Value
INDEX								
30 days/hospital discharge ^d	170/4205 (4.0%)	0.76 (0.50, 0.84)	185/4672 (4.0%)	0.71 (0.40, 0.84)	0.98 (0.79, 1.21)	0.86	0.92 (0.72, 1.18)	0.53
Three months post- OHCA ^e	150/4199 (3.6%)	0.80 (0.67, 0.91)	144/4638 (3.1%)	0.81 (0.68, 1.0)	0.86 (0.68, 1.09)	0.22	1.07 (0.83, 1.38)	0.63
Six months post-OHCA ^f	155/4213 (3.7%)	0.84 (0.70, 1.0)	153/4657 (3.3%)	0.84 (0.67, 1.0)	0.89 (0.70, 1.13)	0.33	0.92 (0.74, 1.15)	0.47
VISUAL ANALOGUE SC.	ALE							
30 days/hospital discharge ^g	173/4208 (4.1%)	70 (50, 80)	182/4669 (3.9%)	65 (45, 80)	0.95 (0.76, 1.17)	0.63	0.81 (0.64, 1.03)	0.08
Three months post- OHCA ^h	152/4201 (3.6%)	80 (60, 90)	145/4639 (3.1%)	80 (65, 90)	0.86 (0.68, 1.08)	0.19	1.08 (0.86, 1.35)	0.53
Six months post-OHCA ⁱ	159/4217 (3.8%)	80 (65, 90)	158/4662 (3.4%)	80 (65, 90)	0.89 (0.70, 1.13)	0.35	1.01 (0.80, 1.27)	0.94

Notes:

^a Outcome is survivors vs non-survivors. Models were adjusted for ambulance service (4 levels), paramedic experience (2 levels: \geq 5 years) and distance from base ambulance station (2 levels: \geq 5 miles). Confidence intervals were adjusted for paramedic clustering using a clustered bootstrap. ^b Outcome is either: (a) EQ-5D single summary index, or (b) EQ-5D visual analogue scale, conditional on surviving to the relevant timepoint. The outcomes were transformed to a scale between 0 and 1 non-inclusive. Models were adjusted for trust (4 levels: YAS, SWAST EMAS and EEAST), paramedic experience (2 levels: \geq 5 years, <5 years) and distance from base ambulance station (2 levels: \geq 5 miles, <5 miles). Confidence intervals were adjusted for paramedic clustering using a clustered bootstrap.

^d Missing for 205 Tracheal Intubation group patients and 214 i-gel group patients.

^e Missing for 211 Tracheal Intubation group patients and 248 i-gel group patients.

^f Missing for 197 Tracheal Intubation group patients and 229 i-gel group patients.

^g Missing for 202 Tracheal Intubation group patients and 217 i-gel group patients.

 $^{\rm h}$ Missing for 209 Tracheal Intubation group patients and 247 i-gel group patients.

ⁱ Missing for 193 Tracheal Intubation group patients and 224 i-gel group patients.





the three timepoints. The 'worst case' and 'imputed case' sensitivity
 analyses, designed to determine the potential impact of missing data,
 did not alter these conclusions.

The majority of RCTs in OHCA have reported only short-term270outcomes, and even the most recent international advisory statement271describing a core outcome set for RCTs in OHCA patients does not272

ARTICLE IN PRESS

RESUSCITATION XXX (2020) XXX -XXX





Eavours TI

cause of the bias.¹² As a described in a previous paper.¹¹ Importantly, the trial population included patients who did and did not receive AAM, and paramedics allocated to the i-gel group were more likely to use an advanced technique than those allocated to TI.

Conclusions

Longer term follow-up confirmed the results of the AIRWAYS-3132 primary analysis. There were no significant differences in functional314outcome or quality of life between the i-gel SGA and TI groups at three315and six months after OHCA. This suggests that our initially reported316findings are robust over time.317

Authors' contribution

Eavours i-gel

319 Guarantors: Benger and Rogers had full access to all the data in the 320 trial and take responsibility for the integrity of the data and the 321 accuracy of the data analysis. Concept and design: Benger, Black, 322 Brett, Kirby, Nolan, Reeves, Robinson, Rogers, Scott, South, Taylor, 323 Thomas, Voss, Wordsworth. Acquisition, analysis, and interpretation 324 of data: All authors. Drafting of the manuscript: Benger. Critical 325 revision of the manuscript for important intellectual content: All 326 authors. Statistical analysis: Lazaroo, Scott, Smartt, Rogers (Clinical 327 Trials and Evaluation Unit, Bristol Medical School, University of 328 Bristol, Bristol, UK). Technical and project support: Clout, Taylor.

Conflict of interest

329

307

308

309

310

311

312

318

All authors have completed and submitted the ICMJE Form for Disclosure of Potential Conflicts of Interest. Rogers salary was funded 331

recommend data collection beyond 90 days, mainly because of the
substantial resources required and the risk of attrition bias.¹² As a
result, the natural history of survivor recovery following OHCA has
been documented by only a few investigators,^{19,22–24} and there
remains a need to examine the longer-term impacts of OHCA on
functional status, cognition and quality of life.^{25,26}

279 Several studies have documented improvements in the functional 280 status of OHCA survivors for at least the first three months and up to six 281 months after cardiac arrest.^{24,25} Our data support this: we have shown 282 a shift in the distribution of mRS scores consistent with improving 283 functional status between hospital discharge and three months, and 284 an attenuated shift in the same direction between three and six 285 months. The decrease in the proportion of patients with an mRS score 286 of 5 between hospital discharge and three months represents a 287 combination of some patients dying and others improving their 288 functional status.27

Although the PART trial documented a significantly higher rate of
 favourable outcome among patients randomised to a strategy of initial
 laryngeal tube SGA compared with TI,¹⁰ longer-term outcomes were
 not collected, so it is unknown whether this difference was sustained at
 three and six months after OHCA.

294 Our research has several limitations. In keeping with similar 295 studies, our trial has relatively few survivors from which to gather 296 longer-term outcomes. Furthermore, we were reliant on both active 297 patient consent and the participant's willingness to complete and 298 return the questionnaires at the follow-up timepoints. Despite 299 considerable effort by the research teams, only 52.4% of survivors 300 consented to active follow up and only 41.3% of survivors were 301 followed up to six months. As a result, our analyses are undermined by 302 missing data, with limited trial power and the risk of attrition bias. 303 However, the proportion of missing data is very similar in the two 304 groups, and there is no evidence that the availability of follow-up data 305 was influenced by patient allocation. Furthermore, the sensitivity 306 analyses reported did not alter our conclusions about the two

8

ARTICLE IN PRESS

RESUSCITATION XXX (2020) XXX -XXX

- by a grant from the British Heart Foundation until March 2017; part of
 Reeves salary was funded by grants from the National Institute for
- Health Research. All other authors declare no conflicts of interest.

335 Funding statement

336 The trial was funded by the National Institute for Health Research 337 (NIHR) Health Technology Assessment (HTA) Programme (project 338 number12/167/102) and supported by the NIHR Comprehensive 339 Research Networks. Professor Benger is a NIHR Senior Investigator. 340 The trial was not funded by any commercial organizations or 341 equipment manufacturers. The views and opinions expressed in this 342 report are those of the authors and do not necessarily reflect those of 343 the HTA, NIHR, NHS or the Department of Health and Social Care.

The funding organization had no role in the design and conduct of
 the trial; collection, management, analysis, and interpretation of the
 data; preparation, review, or approval of the manuscript; and decision
 to submit the manuscript for publication.

This trial was designed and delivered in collaboration with the
 Clinical Trials and Evaluation Unit, a UKCRC registered clinical trials
 unit which, as part of the Bristol Trials Centre, is in receipt of National
 Institute for Health Research CTU support funding.

352 Appendix A. Supplementary data

³⁵³ Supplementary data associated with this article can be found, in the ³⁵⁴ online version, at https://doi.org/10.1016/j.resuscitation.2020.09.026.

REFERENCES

360

362

363

364

366

367

- Hawkes C, Booth S, Ji C, et al. Epidemiology and outcomes from outof-hospital cardiac arrests in England. Resuscitation 2017;110:133 -40.
- Deakin CD. The chain of survival: not all links are equal. Resuscitation 2018;126:80–2.
- Jentzer JC, Clements CM, Wright S, White RD, Jaffe AS. Improving survival from cardiac arrest: a review of contemporary practice and challenges. Ann Emerg Med 2016;68:678–89.
- Gwinnutt CL. Should we intubate patients during cardiopulmonary resuscitation? BMJ 2017;357:j1772.
- 369
 5. Soar J, Nolan JP. Airway management in cardiopulmonary resuscitation. Curr Opin Crit Care 2013;19:181–7.
- Fouche PF, Simpson PM, Bendall J, Thomas RE, Cone DC, Doi SA.
 Airways in out-of-hospital cardiac arrest: systematic review and metaanalysis. Prehosp Emerg Care 2014;18:244–56.
- 376
 7. Sakurai A, Kinoshita K, Maeda Y, et al. Confirmed cardiac output on emergency medical services arrival as confounding by indication: an observational study of prehospital airway management in patients with out-of-hospital cardiac arrest. Emerg Med J 2019;36:410–5.
- Andersen LW, Grossestreuer AV, Donnino AW. Resuscitation time bias – a unique challenge for observational cardiac arrest research. Resuscitation 2018;125:79–82.
- Benger J, Coates D, Davies S, et al. Randomised comparison of the effectiveness of the laryngeal mask airway supreme, i-gel and current

practice in the initial airway management of out of hospital cardiac arrest: a feasibility study. Br J Anaesth 2016;116:262–8.

387

388

380

391

392

393

394

396

397

398

400

401

402

403

405

406

407

408

400

411

413

414

416

417

418

420

421

422

423

425

426

428

429

430

432 433

434

436

437

438

440

442

443

444

446

447

448

459

451

452

454

455

456

458

459

460

- Wang HE, Schmicker RH, Daya MR, et al. Effect of a strategy of initial laryngeal tube insertion vs endotracheal intubation on 72-h survival in adults with out-of-hospital cardiac arrest. A randomized clinical trial. JAMA 2018;320:769–78.
- Benger JR, Kirby K, Black S, et al. Effect of a strategy of a supraglottic airway device vs tracheal intubation during out-of hospital cardiac arrest on functional outcome: the AIRWAYS-2 randomized clinical trial. JAMA 2018;320:779–91.
- Haywood K, Whitehead L, Nadkarni VM, et al. COSCA (Core Outcome Set for Cardiac Arrest) in adults: an advisory statement from the International Liaison Committee on Resuscitation. Circulation 2018;137:e783–801.
- 13. Taylor J, Black S, Brett S, et al. Design and implementation of the AIRWAYS-2 trial: a multi-centre cluster randomised controlled trial of the clinical and cost effectiveness of the i-gel supraglottic airway device versus tracheal intubation in the initial airway management of out of hospital cardiac arrest. Resuscitation 2016;109:25–32.
- Robinson MJ, Taylor J, Brett SJ, et al. Design and implementation of a large and complex trial in emergency medical services. Trials 2019;20:108.
- Joint Royal Colleges Ambulance Liaison Committee. Clinical practice guidelines 2016. London: Class Publishing; 2016.
- Duckett J, Fell P, Han K, Kimber C, Taylor C. Introduction of the i-gel supraglottic airway device for prehospital airway management in a UK ambulance service. Emerg Med J 2014;31:505–7.
- 17. Monsieurs KG, Nolan JP, Bossaert LL, et al. European Resuscitation Council guidelines for resuscitation 2015. Section 1. Executive summary. Resuscitation 2015;95:1–80.
- Elliott VJ, Rodgers DL, Brett SJ. Systematic review of quality of life and other patient-centred outcomes after cardiac arrest survival. Resuscitation 2011;82:247–56.
- 19. Smith K, Andrew E, Lijovic M, Nehme Z, Bernard S. Quality of life and functional outcomes 12 months after out-of-hospital cardiac arrest. Circulation 2015;131:174–81.
- Van Hout B, Janssen MF, Feng YS, et al. Interim scoring for the EQ-5D-5L: mapping the EQ-5D-5L to EQ-5D-3L value sets. Value Health 2012;15:708–15.
- Smithson M, Verkuilen J. A better lemon squeezer? Maximumlikelihood regression with beta-distributed dependent variables. Psychol Methods 2006;11:54–71.
- 22. Andrew E, Nehme Z, Wolfe R, et al. Long-term survival following outof-hospital cardiac arrest. Heart 2017;103:1104–10.
- Lilja G, Nielsen N, Friberg H, et al. Cognitive function in survivors of outof-hospital cardiac arrest after target temperature management at 33°C versus 36°C. Circulation 2015;131:1340–9.
- Steinbuscha CVM, van Heugten CM, Rasquin SMC, et al. Cognitive impairments and subjective cognitive complaints after survival of cardiac arrest: a prospective longitudinal cohort study. Resuscitation 2017;120:132–7.
- Tong JT, Eyngorn I, Mlynash M, Albers GW, Hirsch KG. Functional neurologic outcomes change over the first 6 months after cardiac arrest. Crit Care Med 2016;44:e1202–7.
- Lim C, Verfaellie M, Schnyer D, Lafleche G, Alexander MP. Recovery, long-term cognitive outcome and quality of life following out-of-hospital cardiac arrest. J Rehab Med 2014;46:691–7.
- Arrich J, Zeiner A, Sterz F, et al. Factors associated with a change in functional outcome between one month and six months after cardiac arrest: a retrospective cohort study. Resuscitation 2009;80:876–80.